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EDClust: An EM-MM hybrid method for cell clustering in
population-level single cell RNA sequencing

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Abstract

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Single-cell RNA sequencing (scRNA-seq) technology has revolutionized the genomics research by enabling the measurement of the transcriptomic profile at the level of single cells. One of the most fundamental problems in scRNA-seq data analysis is cell clustering, for which a rather large number of methods have been developed. With the increasing application of scRNA-seq in larger scale studies, people face the problem of cell clustering when the scRNA-seq data are from more than one subject. One challenge in analyzing such data is the subject-specific systematic variations: heterogeneity from multiple subjects may have a significant impact on the clustering accuracy. However, existing methods addressing such effect suffered from several limitations. In this work, we develop a novel statistical method named ‘EDClust’ for scRNA-seq cell clustering when data are from multiple subjects. EDClust models the sequence read counts by a mixture of Dirichlet-Multinomial distributions, and explicitly accounts for the cell type heterogeneity, subject heterogeneity, and the clustering uncertainty. An EM-MM hybrid algorithm is derived for maximizing the data likelihood and clustering the cells. We perform a series of simulation studies to evaluate the proposed method and demonstrate the outstanding performance of EDClust. Comprehensive benchmarking on four real scRNA-seq datasets with various tissue types and species demonstrates the substantial accuracy improvement of EDClust compared to the existing methods.

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1 . Introduction

Single-cell RNA-sequencing (scRNA-seq) is a powerful technology for measuring the gene expression at the single cell level. It offers unprecedented opportunities to answer questions related to cell-specific changes in transcriptome, such as identification of rare cell types and heterogeneity of cell responses¹. Several experimental protocols of scRNA-seq have been developed in the past few years, including SMART-seq², CEL-seq³ and Drop-seq⁴, providing additional choices to meet diverse research needs. Among all, droplet-based technologies encapsulate each individual cell in a nanoliter droplet together with a bead⁴, and substantially reduced the experimental cost. Moreover, droplet-based methods utilize unique molecular identifiers (UMIs) to eliminate the effects of PCR amplification bias⁵. The good scalability, high efficiency and low cost make droplet-based method the top choice for scRNA-seq experiments in population-scale studies.

In scRNA-seq data analysis, the first step is usually cell clustering. The main purpose of clustering is to group cells by their transcriptomic similarity, and then annotate the groups to cell types based on existing biological knowledge. This is a fundamental step in scRNA-seq analysis, since many downstream analyses, including cellular composition estimation, cell type-specific differential expression, and rare cell type discovery, are all carried out based on the clustering results⁶. Though classic unsupervised clustering methods such as K-means and hierarchical clustering can be applied, in view of the sparse and noisy characteristics of scRNA-seq data, many unsupervised methods customized for scRNA-seq data have been developed and widely used. For example, SC3 combines feature selection and dimension reduction in a consensus clustering framework and has been proven to be a highly robust clustering method⁷. Seurat is another popular method that adopts community-detection to identify similar cells and shows good scalability for large datasets⁸. TSCAN fits a mixture of multivariate normal distributions and uses hierarchical clustering to identify cell clusters⁹. Lastly, observing the needs for clustering large-scale study with thousands to millions of cells, SHARP is developed recently for ultra-fast clustering through a divide-and-conquer strategy¹⁰.

All the aforementioned clustering methods are developed without the consideration of systematic biases in the data, that is, the expressions for a gene from all cells in the same cell type are considered to be identically distributed. However, similar to many other high-throughput technologies, scRNA-seq data also suffers from a number of technical biases. One such bias in population

level study is the subject effect: there could be systematic, subject-specific shift in the gene expression. Thus, the distributions of the gene expression can be different between subjects even within the same cell type. That shift can be induced by different characteristics of the subjects, such as demographics or clinical conditions. Or it can be the result of batch effect when different subjects are profiled at different time/location/lab. It is worth mentioning that the batch effects can be severe in scRNA-seq, since it is exacerbated by the fact that most scRNA-seq protocols require fresh tissue for experiments, thus a randomized experimental design for removing batch effect might become impossible in many cases. Nevertheless, most existing clustering methods do not explicitly address the heterogeneity among multiple subjects. Direct application of those methods on data from population studies can lead to inaccurate clustering results due to correlated measurement errors instead of biological similarities¹¹.

One possible remedy for the problem is to consider the subject effect as batch effect, and correct for that before cell clustering. Several computational methods have been developed for batch effect correction and can be applied before clustering. For example, ComBat and ComBat-seq¹² are developed originally for bulk sequencing data and use linear models to remove batch effects. Mutual-nearest-neighbor (MNN) corrects batch effects by constructing a shared space between datasets¹³. Harmony is another popular batch correction method and uses an iterative approach to eliminate batch effects for cells calculated in PCA space¹⁴.

Though it is possible to cluster the cells after removing the subject effects, this “two-step” approach has some drawbacks. First, the batch effect correction procedure often produces negative values for gene expressions, which will generate errors in many cell clustering tools. Secondly, such approach is in general not efficient due to the transformation of data and alteration of data structure. For example, several clustering methods make distributional assumptions on the count data, while the data after batch effect correction is not counts anymore. Such discrepancy will lead to undesirable clustering performances for those methods.

In comparison, a more rigorous and potentially better approach is to design a clustering method that takes subject effects into consideration. Both BAMM-SC¹⁵ and BUSseq¹⁶ are tailored methods for addressing subject effect during clustering. BAMM-SC implements a Bayesian mixture model which utilizes information across genes and individuals to account for the heterogeneity. BUSseq adopts a more complicated hierarchical model that strictly follows the data generation process of

scRNA-seq experiments to correct batch effects and cluster cells. Both methods use Markov Chain Monte Carlo (MCMC) to solve the model, which do not scale well for large datasets. To provide a complementary approach to address the cell clustering problem in population-scale scRNA-seq data, we design EDClust, which is an EM¹⁷ and MM¹⁸ hybrid method based on Dirichlet-Multinomial mixture model, for clustering. EDClust takes the raw count data from multiple subjects without transformation, avoiding the possible destruction of data structure and loss of information. Meanwhile, EDClust explicitly quantifies the effects of heterogeneity from different sources and provides posterior probabilities for cells being in each cluster. Through extensive simulation studies and four real datasets, we show EDClust has better clustering accuracy compared with existing methods. In the following sections, we first introduce the data model and derivation of the EM-MM method. The simulation design and results are presented in Section 3. Lastly, we showcase the performance and utility of EDClust using four real scRNA-seq datasets in Section 4.

2. Methods

To cluster population-scale scRNA-seq data, we propose the following Dirichlet-Multinomial mixture model to capture the cell type and subject effects on gene expression. We assume the number of cell types in the data, K , is known, and all subjects share the same K cell types. Our aim is to cluster all the cells from these subjects simultaneously. Note that K can be specified by investigators based on biological knowledge, or can be determined by a number of software tools. Throughout this work, we will just assume K is known.

2.1 Data model

Let y_{lji} represent the sequence counts for gene j in cell i from subject l ($1 \leq i \leq I_l$, $1 \leq j \leq J$, $1 \leq l \leq L$), where I_l , J and L indicate the total number of cells (in subject l), genes and subjects, respectively. Based on the assumption that $\mathbf{Y}_{li} = (Y_{1li}, Y_{2li}, \dots, Y_{Jli})$ follows a Dirichlet-Multinomial mixture distribution, \mathbf{Y}_{li} can be viewed as generated in two steps. First, a cell type label $W_{li} \in \{1, 2, \dots, K\}$ is assigned to cell i in subject l with probability $\Pr(W_{li} = k) = \pi_{lk}$. Second, given the cell label (i.e., $W_{li} = k$), \mathbf{Y}_{li} will be generated from a Multinomial distribution by $\mathbf{Y}_{li} \sim \text{Multinomial}(T_{li}, p_{li})$. Here, $T_{li} = \sum_j Y_{lji}$ indicates total read counts, and the proportion p_{li} represents the relative gene ex-

pressions. We further assume that p_{li} follows a cell-type specific prior distribution $\text{Dirichlet}(\alpha_{lk}) = \text{Dirichlet}(\alpha_{lk1}, \alpha_{lk2}, \dots, \alpha_{lkJ})$. To simultaneously account for cell type and subject effects, we assume the overall effect α_{lk} can be expressed as the sum of cell type effect α_{0kj} and subject effect δ_{lkj} : $\alpha_{lkj} = \alpha_{0kj} + \delta_{lkj} > 0$. Finally, we assume that all cells in all L subjects are independent and treat cell type label $W_{li} = k$ as the missing data. Then the observed and complete data log likelihood can be written as:

$$l(\mathbf{Y}; \mathbf{T}, \Theta) = \sum_{l=1}^L \sum_{i=1}^{I_l} \log \left[\sum_{k=1}^K \pi_{lk} P(Y_{li} | T_{li}, \alpha_{0k} + \delta_{lk}) \right] \quad (2.1)$$

$$l_c(\mathbf{Y}, \mathbf{W}; \mathbf{T}, \Theta) = \sum_{l=1}^L \sum_{i=1}^{I_l} \sum_{k=1}^K I(W_{li} = k) [\log \pi_{lk} + \log P(Y_{li} | T_{li}, \alpha_{0k} + \delta_{lk})] \quad (2.2)$$

Here $\mathbf{W} = \{W_{li} : i = 1, \dots, I_l, l = 1, \dots, L\}$ includes the indicator of cell type labels, and $\Theta = \{\pi_{lk}, \alpha_{0k}, \delta_{lk} : k = 1, \dots, K, l = 1, \dots, L\}$ contains all the model parameters. $P(Y_{li} | T_{li}, \alpha_{0k} + \delta_{lk})$ represents the Dirichlet-Multinomial probability density, which is

$$P(Y_{li} | T_{li}, \alpha_{0k} + \delta_{lk}) = \binom{T_{li}}{Y_{li}} \frac{\prod_{j=1}^J (\alpha_{0kj} + \delta_{lkj}) (\alpha_{0kj} + \delta_{lkj} + 1) \cdots (\alpha_{0kj} + \delta_{lkj} + Y_{lij} - 1)}{\|\alpha_{0k} + \delta_{lk}\|_1 (\|\alpha_{0k} + \delta_{lk}\|_1 + 1) \cdots (\|\alpha_{0k} + \delta_{lk}\|_1 + T_{li} - 1)} \quad (2.3)$$

Here, $\|\alpha_{0k} + \delta_{lk}\|_1 = \sum_{j=1}^J |\alpha_{0kj} + \delta_{lkj}| = \sum_{j=1}^J (\alpha_{0kj} + \delta_{lkj})$.

2.2 The EM-MM hybrid algorithm for maximum likelihood

The introduction of the latent variable \mathbf{W} allows one to implement the EM algorithm to maximize the observed data likelihood and obtain posterior probabilities for cell type assignment (W_{li}). An EM algorithm iterates between two steps: an expectation step (E-step) and a maximization step (M-step)¹⁷. Let $\Theta^{(t)}$ be the parameter estimate in iteration t . In the E-step, we compute the conditional expectation of W_{li} :

$$\begin{aligned}\mu_{lik}^{(t)} &= E[I(W_{li} = k)|\mathbf{Y}, \Theta^{(t)}] = P(W_{li} = k|\mathbf{Y}, \Theta^{(t)}) \\ &= \frac{\pi_{lk}^{(t)} P(Y_{li}|T_{li}, \alpha_{0k}^{(t)} + \delta_{lk}^{(t)})}{\sum_{k'=1}^K \pi_{lk'}^{(t)} P(Y_{li}|T_{li}, \alpha_{0k'}^{(t)} + \delta_{lk'}^{(t)})}\end{aligned}\quad (2.4)$$

In the M-step, we maximize the ‘‘Q function’’ (the expected complete data log-likelihood with respect to Θ) to obtain $\Theta^{(t+1)}$. The update for π_{lk} can be obtained by solving $\partial Q/\partial \pi_{lk} = 0$.

$$Q(\Theta|\Theta^{(t)}) = E[l(\Theta)|\mathbf{Y}, \Theta^{(t)}] = \sum_{l=1}^L \sum_{i=1}^{I_l} \sum_{k=1}^K \mu_{lik}^{(t)} [\log \pi_{lk} + \log P(Y_{li}|T_{li}, \alpha_{0k} + \delta_{lk})] \quad (2.5)$$

$$\pi_{lk}^{(t+1)} = \frac{\sum_{i=1}^{I_l} \mu_{lik}^{(t)}}{I_l} \quad (2.6)$$

The M-step derivation for α_{0k} and δ_{lk} is much more difficult, and there is not closed form solution. For that, we design the following Minorization-Maximization (MM) algorithm¹⁸ for updating α_{0k} and δ_{lk} . Conceptually, to maximize an objective function $f(\theta)$, an MM algorithm iterates between two steps. In the first step, one uses the current parameter estimate $\theta^{(n)}$ to construct a surrogate function $g(\theta|\theta^{(n)})$ such that $g(\theta|\theta^{(n)})$ minorizes $f(\theta)$, i.e.,

$$\begin{aligned}f(\theta) &\geq g(\theta|\theta^{(n)}) \quad \forall \theta \neq \theta^{(n)} \\ f(\theta^{(n)}) &= g(\theta^{(n)}|\theta^{(n)})\end{aligned}\quad (2.7)$$

In the second step, one finds θ to maximize the surrogate function $g(\theta|\theta^{(n)})$, which gives a new parameter estimate $\theta^{(n+1)}$. Since

$$f(\theta^{(n+1)}) \geq g(\theta^{(n+1)}|\theta^{(n)}) \geq g(\theta^{(n)}|\theta^{(n)}) = f(\theta^{(n)}) \quad (2.8)$$

$f(\theta^{(n)})$ will never decrease as n increases. The algorithm will converge to a stationary point, usually a mode of the objective function.

Extending the work by Zhou and Lange¹⁹, we rewrite the log-likelihood function in (2.2) as the following.

$$\begin{aligned}
I(\Theta) &= \sum_{l=1}^L \sum_{k=1}^K \left[\sum_{i=1}^{I_l} I(W_{li} = k) \log \pi_{lk} - \sum_{c_{1l}} r_{lkc} \log(\|\alpha_{0k} + \delta_{lk}\|_1 + c_{1l}) + \right. \\
&\quad \left. \sum_{j=1}^J \sum_{c_{2lj}} s_{lkjc} \log(\alpha_{0kj} + \delta_{lkj} + c_{2lj}) \right] + \text{const.} \quad (2.9) \\
r_{lkc} &= \sum_{i=1}^{I_l} I(W_{li} = k) I(T_{li} \geq c_{1l} + 1), \quad s_{lkjc} = \sum_{i=1}^{I_l} I(W_{li} = k) I(Y_{lji} \geq c_{2lj} + 1)
\end{aligned}$$

Here the index c_{1l} ranges from 0 to $\max_i(T_{li}) - 1$, and the index c_{2lj} runs from 0 to $\max_i(Y_{lji}) - 1$. In MM algorithm, we design a surrogate function that minorizes the log-likelihood function. Assuming that $\alpha_{0kj} > 0$ and $\delta_{lkj} \geq 0$, we can utilize the following inequalities:

$$-\log(c + \|\alpha_{lk}\|_1) \geq -\frac{1}{\|\alpha_{0k}^{(n)} + \delta_{lk}^{(n)}\|_1 + c} (\|\alpha_{0k} + \delta_{lk}\|_1) + \text{const.} \quad (2.10)$$

$$\log(\alpha_{0kj} + \delta_{lkj} + c) \geq \frac{\alpha_{0kj}^{(n)}}{\alpha_{0kj}^{(n)} + \delta_{lkj}^{(n)} + c} \log(\alpha_{0kj}) + \frac{\delta_{lkj}^{(n)}}{\alpha_{0kj}^{(n)} + \delta_{lkj}^{(n)} + c} \log(\delta_{lkj}) + \text{const.} \quad (2.11)$$

For them, the equality holds when $\alpha_{0kj} = \alpha_{0kj}^{(n)}$ and $\delta_{lkj} = \delta_{lkj}^{(n)}$. We construct the following surrogate function $g(\Theta|\Theta^{(t,n)})$ as:

$$\begin{aligned}
g(\Theta|\Theta^{(t,n)}) &= \sum_{l=1}^L \sum_{k=1}^K \left\{ \sum_{i=1}^{I_l} \mu_{lik}^{(t)} \log \pi_{lk} - \sum_{c_{1l}} r_{lkc}^{(t)} \frac{\|\alpha_{0k} + \delta_{lk}\|_1}{\|\alpha_{0k}^{(t,n)} + \delta_{lk}^{(t,n)}\|_1 + c_{1l}} + \right. \\
&\quad \left. \sum_{j=1}^J \sum_{c_{2lj}} s_{lkjc}^{(t)} \left[\frac{\alpha_{0kj}^{(t,n)} \log(\alpha_{0kj})}{\alpha_{0kj}^{(t,n)} + \delta_{lkj}^{(t,n)} + c_{2lj}} + \frac{\delta_{lkj}^{(t,n)} \log(\delta_{lkj})}{\alpha_{0kj}^{(t,n)} + \delta_{lkj}^{(t,n)} + c_{2lj}} \right] \right\} + \text{const.} \quad (2.12) \\
r_{lkc}^{(t)} &= \sum_{i=1}^{I_l} \mu_{lik}^{(t)} I(T_{li} \geq c_{1l} + 1), \quad s_{lkjc}^{(t)} = \sum_{i=1}^{I_l} \mu_{lik}^{(t)} I(Y_{lji} \geq c_{2lj} + 1)
\end{aligned}$$

By solving $\partial g(\Theta|\Theta^{(t,n)})/\partial \delta_{lkj} = 0$ and $\partial g(\Theta|\Theta^{(t,n)})/\partial \alpha_{0kj} = 0$, we obtain the MM updates for δ_{lkj} and α_{0kj} as:

$$\delta_{lkj}^{(t,n+1)} = \left(\sum_{c_{2lj}} \frac{s_{lkjc}^{(t)} \delta_{lkj}^{(t,n)}}{c_{2lj} \alpha_{0kj}^{(t,n)} + \delta_{lkj}^{(t,n)} + c_{2lj}} \right) / \left(\sum_{c_{1l}} \frac{r_{lkc}^{(t)}}{\|\alpha_{0k}^{(t,n)} + \delta_{lk}^{(t,n)}\|_1 + c_{1l}} \right) \quad (2.13)$$

$$\alpha_{0kj}^{(t,n+1)} = \left(\sum_{l=1}^L \sum_{c_{2lj}} \frac{s_{lkjc}^{(t)} \alpha_{0kj}^{(t,n)}}{c_{2lj} \alpha_{0kj}^{(t,n)} + \delta_{lkj}^{(t,n)} + c_{2lj}} \right) / \left(\sum_{l=1}^L \sum_{c_{1l}} \frac{r_{lkc}^{(t)}}{\|\alpha_{0k}^{(t,n)} + \delta_{lk}^{(t,n)}\|_1 + c_{1l}} \right) \quad (2.14)$$

Within the M-step in each EM iteration, EDClust runs multiple MM iterations to update α_0 and δ . To reduce the computational burden, we only run 3 MM iterations in each M-step. Real data analyses show that such procedure provides comparable performance as running more (such as 20) iterations.

2.3 Feature selection

Feature selection is one of the key steps before clustering. We aim to select a subset of informative genes that can identify the structure of data and thus improve the performance of clustering. A recently developed feature selection tool tailored to scRNA-seq, FEature SelecTion (FEAST)²⁰, shows great potential for improving clustering accuracy. FEAST computes the F-statistics for each feature based on embedded consensus clustering results and provides a ranking list of feature significant. By default, EDClust applies FEAST to obtain the top 500 features for clustering. In the software implementation, users have option to specify the gene features.

2.4 Determine the initial values

It is known that EM algorithm often suffers from locally optimal solutions. Our problem, due to the high dimensionality and complex nature of the data, is particularly prone to such problem. Thus, it is crucial to provide good initial values for the parameters, especially α_0 and δ . For that, we run unsupervised clustering on a randomly chosen subject to obtain initial clusters, and then the initial value for α_0 can be computed from these initial clusters. To be specific, we set a randomly chosen subject as the baseline, and thus its overall effect is entirely contributed by the cell type effect α_0 . Based on the initial clusters, we obtain a naive estimate $\hat{\alpha}_0$ according to the relative gene expression in each initial cluster and take it as the initial value. We set the selected subject with a subject effect of zero, and set initial values for the rest of δ 's to be small positive numbers (10^{-5} by default).

We use SHARP¹⁰ as the unsupervised clustering method to determine initial values due to its good computational performance.

2.5 Software implementation

Overall, the complete EDClust algorithm is summarized in Figure 1. EDClust is designed as an R package with core algorithm written in Julia for better computational efficiency. Scripts for EDClust are available at <https://github.com/weix21/EDClust>. We are working to develop a software package and will submit to Bioconductor soon.

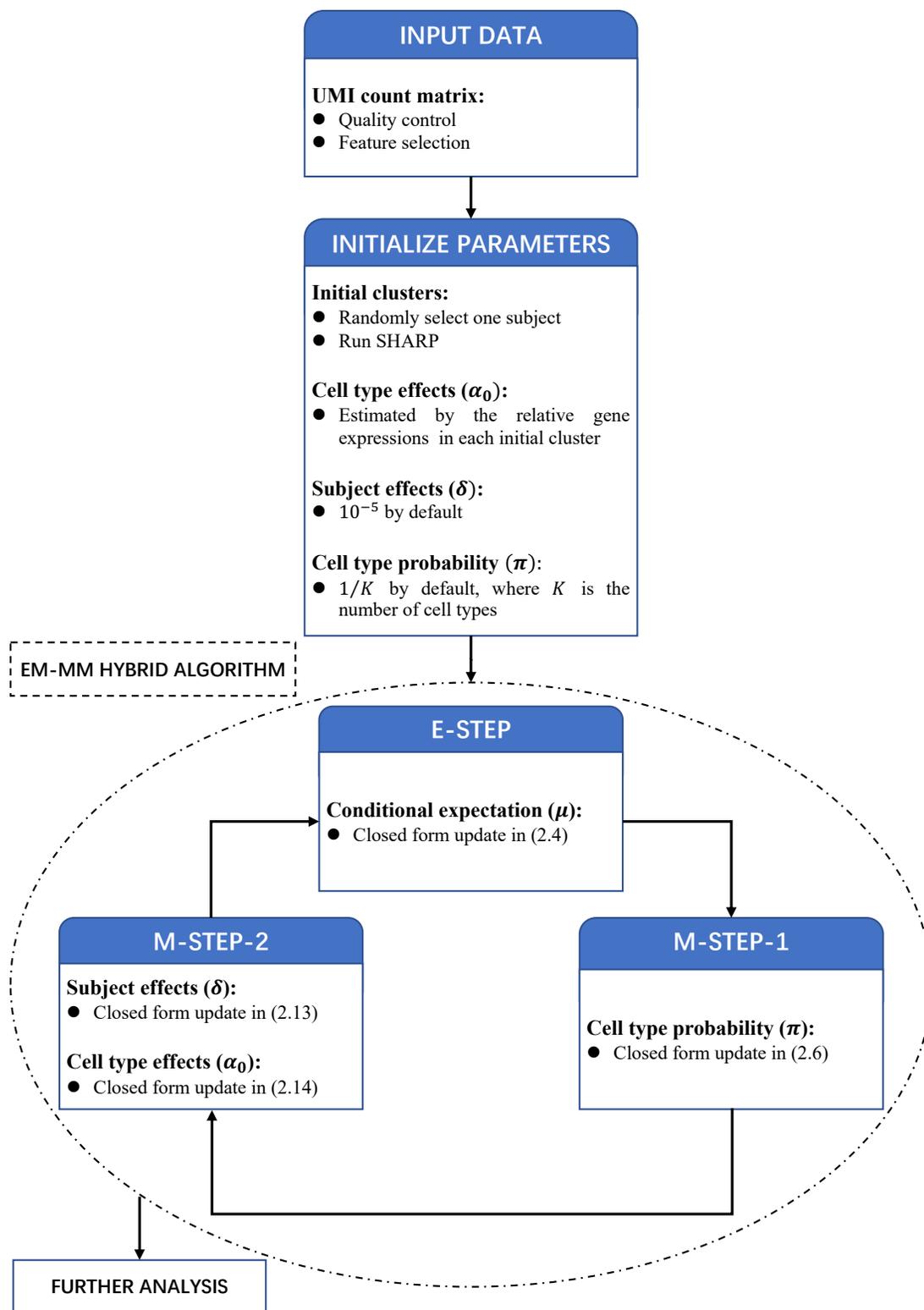


Figure 1: Summary of the EDClust algorithm.

3. Simulation studies

We design a series of simulation studies to comprehensively evaluate the performance of EDClust, and compare it to a number of competing methods. The simulations are based on a set of real scRNA-seq from human skin studies (described later). We evaluate the methods when data have different levels of subject specific effects (low, medium and high), and with different sample size selections (5, 10, 15). Specifically, for gene j , cell i and subject l , we generate observed single cell RNA-seq counts from a Dirichlet-Multinomial distribution by $Y_{li} \sim \text{Multinomial}(T_{li}, p_{li})$ where $p_{li} \sim \text{Dirichlet}(\alpha_{lk})$. The total read counts T_{li} match the real data. We specify a log-normal prior distribution on α_{lk} as $\alpha_{lk} \sim \text{LogNormal}(\mu_k, \sigma_k^2)$. The mean μ_k and variance σ_k^2 are first pre-defined for each random sample and cell types by adopting hyper-prior $\mu_k \sim \text{Uniform}(0, 1)$ and $\sigma_k^2 \sim \text{Gamma}(\beta, \tau)$. In all the simulations, we specify τ as 1 and changed the value of β to control the cross subject heterogeneity in the data. Larger heterogeneity indicates stronger subject-specific effects, and thus it is more difficult to cluster.

We compare EDClust with the other four unsupervised clustering methods (SC3, Seurat, SHARP and TSCAN) which embed Harmony for batch effect correction before clustering. We use adjusted Rand index (ARI)²¹ as the evaluation criterion to benchmark the predicted cell type labels. All the simulation results are summarized over 20 Monte Carlo data sets.

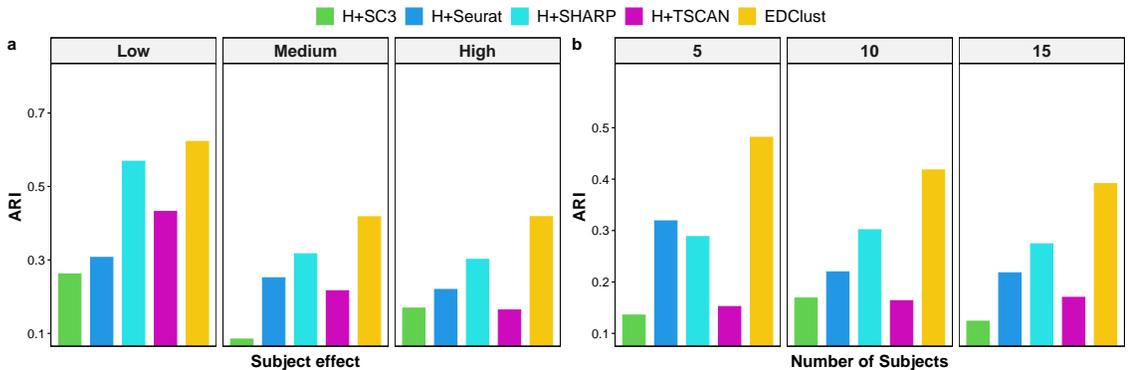


Figure 2: Barplots of average ARI for 5 clustering methods across over 20 simulations, where “H +” indicates that the simulation data are processed by Harmony to remove the subject effects. **a** Influence of subject effect on clustering results. Simulation data is consisted of 10 subjects. **b** Influence of the number of subjects on clustering results with medium level of the subject effect.

As shown in Figure 2a, across three scenarios with different levels of cross subject heterogeneity, EDClust constantly achieves the highest average ARI. The performance of EDClust remains stable

even as the subject effect varies from medium level to high level. Figure 2b presents the influence of number of subjects on clustering. As the number of subjects increased, so did the sources of heterogeneity. As expected, the average ARIs for most of the clustering methods decrease when the number of subjects increases. EDClust consistently outperforms other four methods in terms of ARI. The simulation studies showcase great potential of EDClust in accounting for subject-specific effects and clustering population-scale scRNA-seq data with outstanding performance.

4 . Real data analyses

We benchmark EDClust and other methods on four sets of real scRNA-seq with multiple subjects. More description of the datasets and data processing procedures are provided in each of the subsections below. Here in Table 1, we present the overall results for all four datasets, including the mean and standard deviation of ARIs from 50 runs in each datasets. In addition to the four clustering methods compared in the simulation study, we also compare EDClust with DIMM-SC and BAMM-SC, which have similar model assumptions. Since both TSCAN and Seurat are deterministic clustering methods, they don't have standard deviation in the results. The average ARIs are also displayed in Figure 3. These results show that for three out of the four dataset, EDClust has the best performance, and the performance improvement can be significant. For example in the Mouse Retina data, EDClust has mean ARI 0.87, while the second best performer (Harmony+SC3) only has ARI 0.70. In the Mouse Lung data, EDClust performs slightly worse than BAMM-SC and Harmony+SC3, but not very far off.

Table 1: The ARI of fifty times clustering analyses for each method on four real datasets

Method	Mouse Retina		Baron Pancreas		Human Skin		Mouse Lung	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD
BAMM-SC	0.4273	0.0058	0.6411	0.0686	0.7732	0.0688	0.7354	0.0323
DIMM-SC	0.4221	0.0065	0.6968	0.0692	0.7975	0.0839	0.7003	0.0643
Harmony+SC3	0.6972	0.0687	0.5590	0.1035	0.8260	0.0249	0.7652	0.0026
Harmony+Seurat	0.1024	-	0.5137	-	0.6520	-	0.5307	-
Harmony+SHARP	0.6572	0.0095	0.3115	0.0236	0.8369	0.0533	0.7029	0.0271
Harmony+TSCAN	0.2905	-	0.6392	-	0.6486	-	0.6354	-
EDClust	0.8735	0.0251	0.8047	0.0747	0.9191	0.0813	0.7227	0.0435

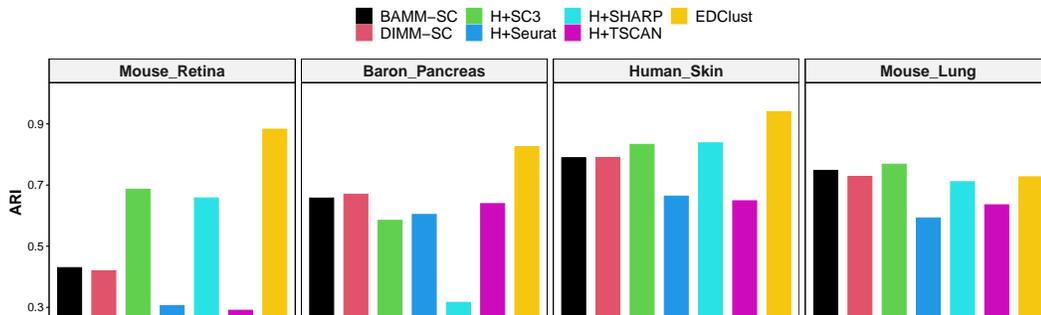


Figure 3: Bar plots of performance of EDClust and competing methods measured by average ARI across 50 clustering results. “H +” represents that the clustering methods are implemented on the real datasets which have been processed by Harmony to remove the batch effects.

4.1 Mouse Retina dataset

We first evaluate the clustering performance of EDClust in mouse tissues through a mouse retina dataset, which is collected from 14-day-old mice in seven batches⁴. Cells are first pooled together to filter out low-expression genes based on dropout rate. We apply FEAST to generate a ranking list of features, and select top 500 genes based on it. Five major cell types are retained and the number of cells is 43,603.

As shown in Table 1 and Figure 3, most of the methods struggle on this dataset with very low average ARI. Though the performances of Harmony+SC3 and Harmony+TSCAN are slightly better, their average ARIs are still below 0.70. EDClust achieves the highest ARI (0.8735), suggesting the excellent performance of EDClust. To visualize the clustering results, we generate some t-SNE plots as shown in Figure 4. The t-SNE plot generated based on the clustering result of EDClust is highly similar to the t-SNE plot with the true labels. We also show the t-SNE plot based on the clustering results from BAMB-SC and Harmony+SC3, where the circled regions highlight the incorrectly clustered cells. These plots provide clear visualization for demonstrating the improved performance of EDClust.

4.2 Baron Pancreas dataset

To evaluate the performance of EDClust in human tissue, we analyze a set of human pancreas data (named “Baron Pancreas” dataset). The original data includes over 12,000 pancreatic cells from four human donors and two mouse strains²². We extract cells from the human donors and filter out

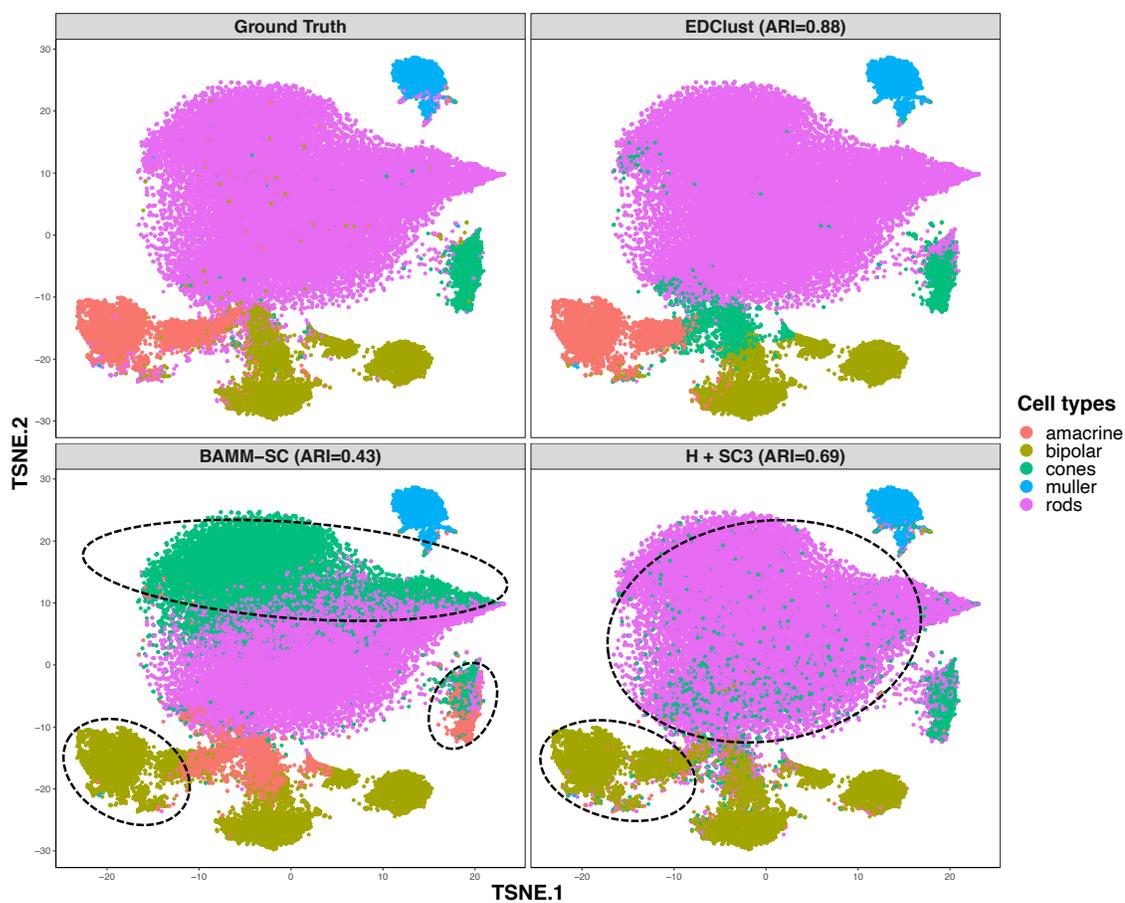


Figure 4: The t-SNE plots of cells in the mouse retina dataset⁴. Each plot is colored by the ground truth, labels inferred by EDClust, BAMB-SC, and Harmony+SC3 (H+SC3), respectively.

the lowly expressed genes. The processed data contains 500 genes a total of 8,506 cells. Some very rare cell types with only a few cells (such as T cells) are removed and 10 major cell types are kept for further analysis.

Both Table 1 and Figure 3 show that the average ARI of EDClust is up to 0.8047 while all other methods fail to achieve the average ARI of 0.70. Figure 5 elucidates that Harmony+SC3 mixed massive cells. Compared to BAMM-SC, EDClust correctly identifies beta cells. Based on EDClust, for most of the cells, we are able to assign labels that are close to the approximated truth. These results showcase the outstanding performance of EDClust in the Baron Pancreas dataset.

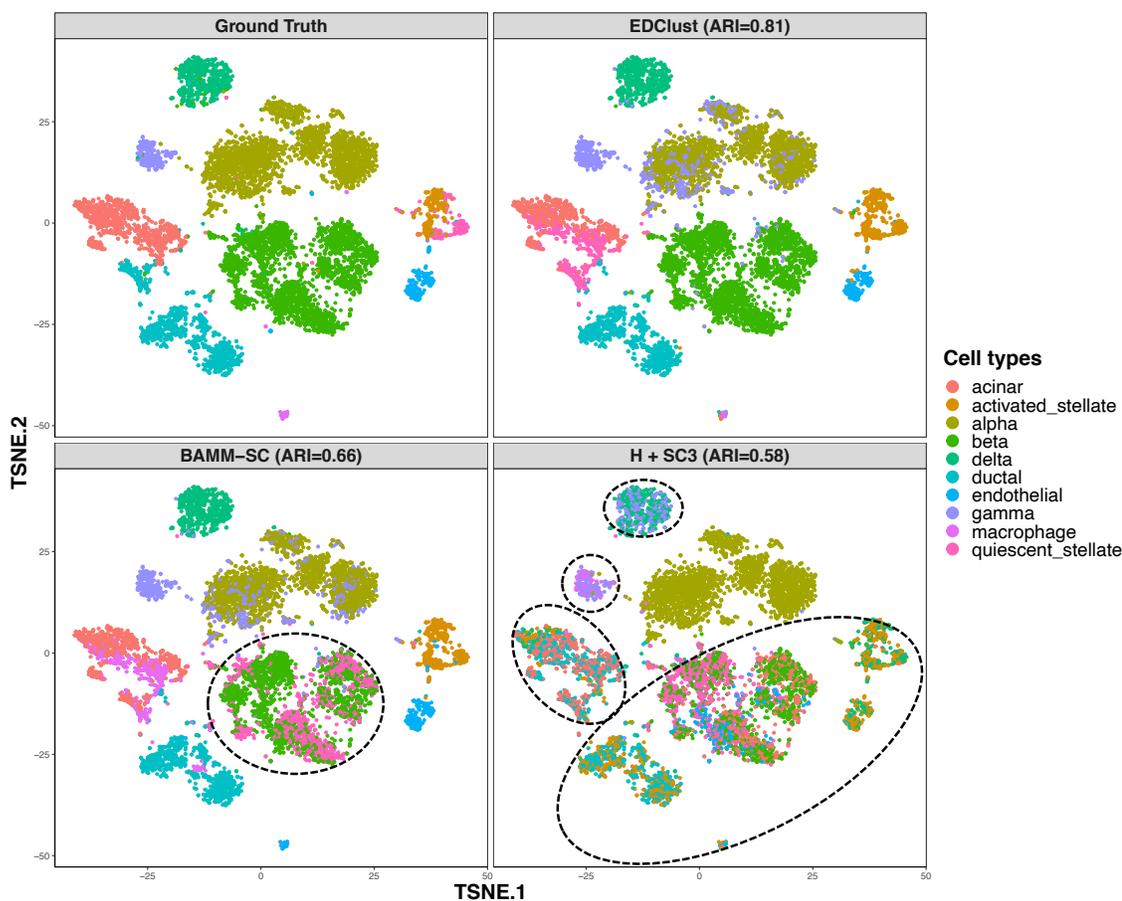


Figure 5: The t-SNE plots of cells in the Baron Pancreas dataset²². Each plot is colored by the ground truth, labels inferred by EDClust, BAMM-SC, and Harmony+SC3 (H+SC3), respectively.

4.3 Human Skin dataset

We further evaluate the clustering performance of EDClust in another human tissues through a human skin dataset, which includes skin samples collected from three health donors in a systemic

sclerosis study¹⁵. In their study, Sun et al.¹⁵ identified eight major types of cells. We use their results as the ground truth, but remove cells with uncertain cell type. After quality control and feature selection, 3,067 cells with 500 selected genes are used in the clustering analysis.

From Table 1 and Figure 3, we can find that EDClust has the most outstanding performance (average ARI = 0.9191) among all the methods. Although the average ARIs for most methods are close to 0.80, EDClust is more accurate in the clustering of several cell types. As shown in Figure 6, compared with BAMB-SC and Harmony+SC3, macrophages/DC, basal keratinocytes and suprabasal keratinocytes can all be classified by EDClust and each is assigned a specific cell type label, which points out the superior performance of EDClust on the Human Skin dataset.

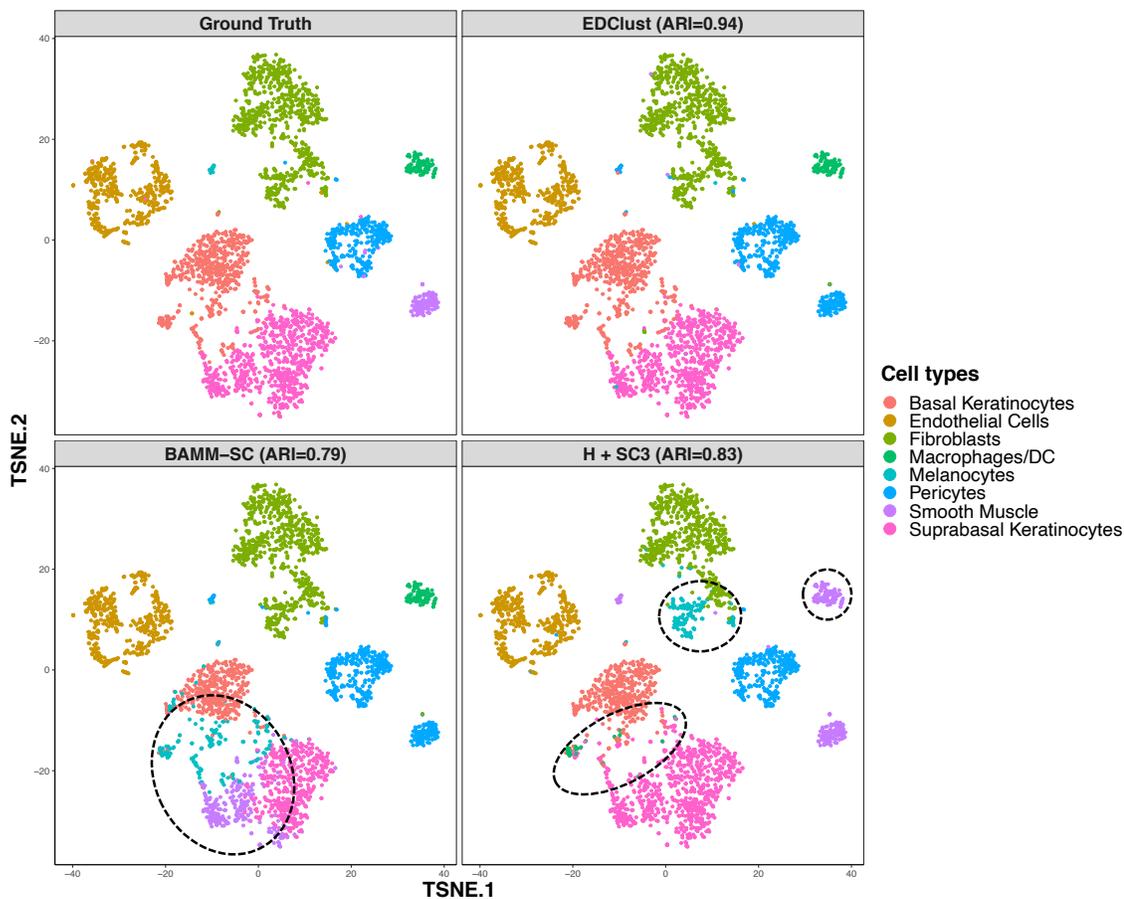


Figure 6: The t-SNE plots of cells in the Human Skin dataset¹⁵. Each plot is colored by the ground truth, labels inferred by EDClust, BAMB-SC, and Harmony+SC3 (H+SC3), respectively.

4.4 Mouse Lung dataset

At last, we evaluate the performance of EDClust in a real dataset with fewer cells. We mainly analyze a mouse lung dataset, which is obtained by collecting lung mononuclear cells from four mouse samples in *Streptococcus pneumoniae* (SP) infected group and control group¹⁵. After data processing step, we obtain 500 top features provided by FEAST and a total of 1,756 cells. Each cell is assigned a cell type label according to previous study by Sun et al.¹⁵ and the expected number of clusters is set as six.

All methods have similar performances on the Mouse Lung dataset (Table 1 and Figure 3). The performance of EDClust (average ARI = 0.7227) is slightly lower than Harmony+SC3 and BAMM-SC. Figure 7 presents consistent pattern. In general, despite some mixed cell types, EDClust has an excellent performance in characterization of endothelial cells and neutrophils.

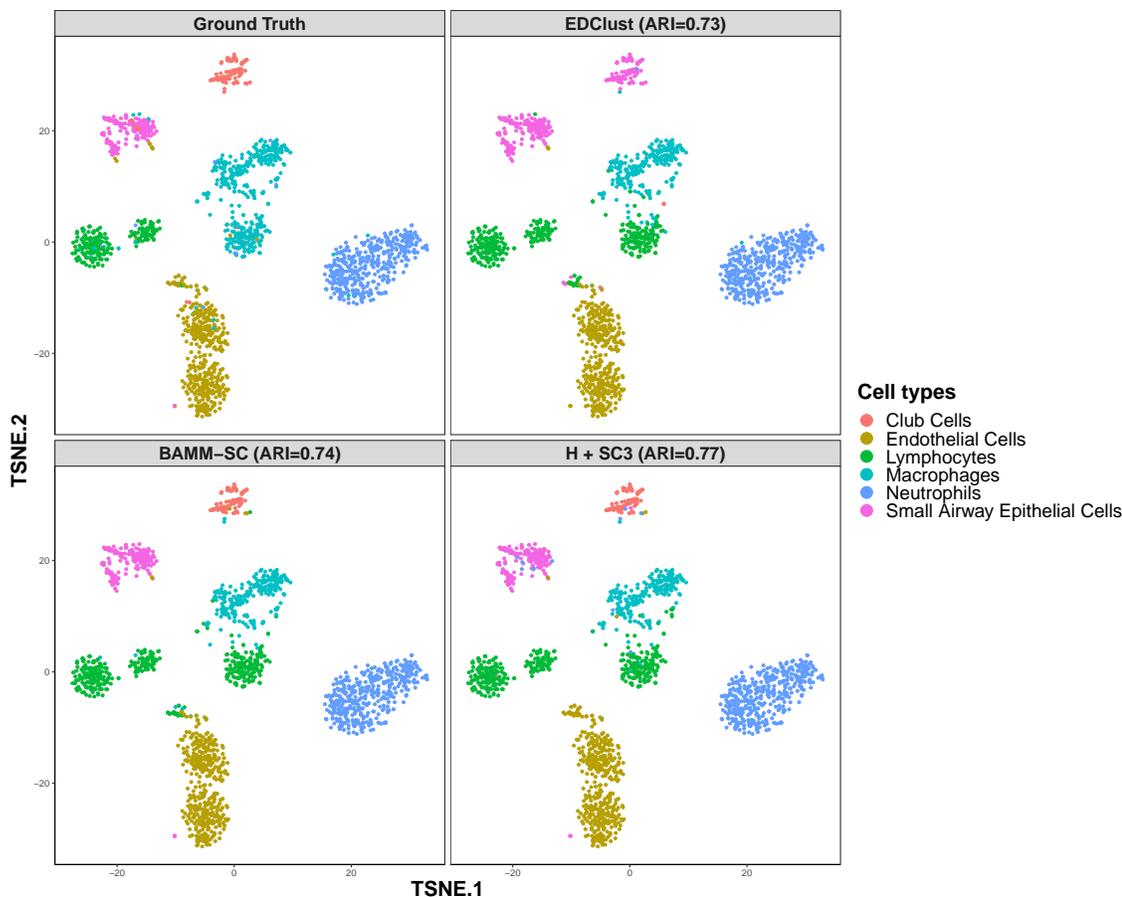


Figure 7: The t-SNE plots of cells in the Mouse Lung dataset¹⁵. Each plot is colored by the ground truth, labels inferred by EDClust, BAMM-SC, and Harmony+SC3 (H+SC3), respectively.

4.5 Computational performance

The EM algorithm usually converges slowly and has heavy computational burden. Our proposed method embeds a few MM iteration within each EM iteration, which brings higher computational cost. However, we implement the software in Julia, with an interface to R, and achieves reasonable computational performance. We benchmark the computational performances of all methods under comparison shown in Table 2. It shows that for the biggest dataset we have tried (Mouse Retina dataset with 7 batches and 43,603 cells), it takes about 35 minutes on a normal computer with a single process. This is considerably faster than BMM-SC, which serves the same purpose but uses MCMC. Other methods either ignore the subject effects or perform a two-step approach (batch effect removal then cell clustering), which are not really comparable to EDClust. We will continue our development of EDClust and implement parallel computing to further improve the performance. Overall, EDClust provides satisfactory computational performance.

Table 2: The average computational time (in min) of each method on four real datasets

Method	Mouse Retina	Baron Pancreas	Human Skin	Mouse Lung
BMM-SC	180	90	13	7
DMM-SC	4	3	2	1
Harmony+SC3	57	24	6	8
Harmony+Seurat	34	4	1	<1
Harmony+SHARP	37	4	2	<1
Harmony+TSCAN	33	4	1	<1
EDClust	35	60	25	10

5. Discussion

In this work, we develop a novel statistical method for cell clustering in multi-subject scRNA-seq data. We model the read counts by a Dirichlet-Multinomial mixture distribution, where the Dirichlet parameters contain subject and cell type effects. We develop an EM-MM hybrid algorithm for fitting the mixture model and performing model-based clustering.

Compared to existing clustering methods that ignore the subject specific effects, EDClust has the following advantages: (1) EDClust provides a tool to describe data heterogeneity among multiple subjects and more effectively identify subject-specific cell types. (2) Utilizing the shared infor-

mation among subjects, EDClust clusters all the cells from all subjects at the same time, which improves the accuracy of cell clustering. (3) Most of the clustering methods can only be performed after several preprocessing approaches, e.g. normalization and batch effect removal, while EDClust offers a one-stop service which can be directly applied on raw count data. (4) EDClust quantifies cluster uncertainty with the probability that each cell belongs to a given cluster, contributing to further statistical inference and biological interpretation.

In our simulation studies, we investigate the influence of heterogeneity among multiple subjects on clustering results. We generate simulated data from Dirichlet-Multinomial distribution with specifying different subject effects and different sample size selections. In real data analyses, we compare the performance of EDClust with competing methods in four droplet-based scRNA-seq datasets with multiple subjects collected from different tissue types or species. In contrast to existing clustering methods that work on data with batch-effect removal through Harmony, EDClust considerably improves clustering accuracy under various experimental designs. Compared to existing clustering methods which account for subject variability, EDClust adopts a straightforward and highly explanatory model but at the same time outperforms those methods in terms of ARI on most of the datasets we tested.

For real data analysis, since the initial values of the cell type effects are set as the naive estimates based on the clustering results given by SHARP, we recommend running EDClust multiple times, each time using a different random seed, and select the one with the best likelihood as the final result. Estimation of $\sum_{j=1}^J \alpha_{0kj}$ provide by Ronning²³ or moment estimates proposed by Weir and Hill²⁴ can also be an appropriate choice for obtaining initial values. Moreover, determination of the number of clusters is a crucial step. We suggest predefining it based on prior biological knowledge or model selection criteria such as AIC²⁵ and BIC²⁶.

There are several limitations of EDClust. First, similar to all the methods using EM algorithm, EDClust could be sensitive to the initial values. The current algorithm computes initial values based on existing unsupervised clustering method. How to better determine initial values is our research direction in the near future. For example, we could use a small subset of cells with strong confidence in cell types, or apply prior biological knowledge. Secondly, EDClust is computationally heavy, especially when the dataset contains some cells with a large total count. We will investigate whether we can build a different surrogate function in order to improve the computational performance.

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